

### Supplementary Figure 1 | Example of ABSOLUTE analysis

**a-e,** An HGS-OvCa sample processed using ABSOLUTE, as in **Figure 1b-g**. In this sample, the excellent SCNA-fit score is strongly opposed to a karyotype score indicating a simpler solution, and dominates the final ranking.



# Supplementary Figure 2 | Karyotype models for GBM, constructed by clustering tumor sets by absolute

**(top)** it of the multiveriate multiveriate multiveriate model (Online Mehtods Eq. 8) using 3 components. Grey-scale denote the probability of each phologous chromosome-arm (y-axis) attaining copy-numbers 0-6 (x-axis). **(bottom)** individual GBM samples are shown grouped according to their cluster membership in (top). Color-scale indicates absolute copy-number (blue - 0, white - 1, red - 2, yellow- 3, green -4. Note - low-copy HSCRs are shown on top, high-numbers on the bottom.





### Supplementary Figure 3 | Unusual karyotypes identified using ABSOLUTE.

a-d, Analysis of homologous copy-ratio profiles with ABSOLUTE, as in Figure 1.

**a**,**b**, Near-haploid genomes

c,d, Hyperaneuploid genomes

**a**, Lung adenocarcinoma sample SM-11ZY. Purity = 0.36, ploidy = 1.12.

**b**, Glioma sample 'glioma 612'. Purity = 0.88, ploidy = 1.14.

**c**, HGS-OvCa sample TCGA-25-1320-01A-01D-0452-01. Purity = 0.58, ploidy = 6.03.



**Supplementary Figure 4 | Identification of a leukocyte methylation signature in HGS-OvCa samples.** Leukocyte methylation signature (Online Methods).



Allelic copy ratio

Supplementary figure 5 | Homologous copy-ratio profiles of called and uncalled cancer samples

**a-d,** Copy profiles for 12 cancer samples. Rows correspond to called samples (**a**), and three distinct failure modes (**b-d**), as indicated (left-hand labels). Homologous copy profiles are colored according to allelic imbalance, as in **Fig. 1c**. The distribution of each outcome is shown for each tumor type in **Fig. 3a**.

a, Called samples annotated with fit homologue-specific copy numbers (HSCN; top-axis).

**b**, The most common failure-mode designation was "non-aberrant" (9.1% of samples), indicating that the copy profile was indistinguishable from that of a normal (diploid) sample.

**c**, An additional 7.3% of samples failed due to insufficient purity. In these samples, copy aberrations were apparent, but so attenuated as to obscure the pattern of integer copy states.

**d**, The remaining 6.9% of failed samples were designated as "polygenomic"; although clear SCNAs were observed, they were not consistent with a single dominant copy profile (Eq. 1).



## Supplementary figure 6 | Distribution of estimated purity in various tumor types, and effect on observed copy-ratios

**a**,**b**, Data are shown for each indicated sample-type, with samples sorted by tumor purity (left). Total copy ratios in were median-centered at 1 for plotting.

**a**, Genome wide copy-ratio profiles are shown for several tumor cohorts. The effect of tumor contamination is evident as the 'fading' of copy ratios towards 1 as tumor purity decreases.

**b**, The data from **a** are shown after '*in silico*' removal of tumor contamination by dividing total absolute copynumber by sample ploidy. This transformation enables more direct comparison between datasets, as samples without copy alterations may be clearly distinguished from heavily contaminated tumors.



### Supplementary figure 7 | Effect of tumor purity and ploidy on power for detection of somatic pointmutations by sequencing

**a**, Combinations of tumor purity and ploidy which imply equal values of  $\delta = \alpha/D$  (Eqs. 1,8), the concentration-ratio of molecules present at a single copy per cancer cell (multiplicity 1) in the tumor population. Dotted lines indicate equal values of  $2(1-\alpha)/D$  (Eq. 1).

**b**, Theoretical power to detect clonal somatic point mutations present at multiplicity 1, as a function of sequence coverage, for various values of  $\delta$  (shown in **a**). Power was calculated for FPR <=  $5 \times 10^{-7}$ , assuming a uniformly random sequencing error-rate of  $10^{-3}$  (Online Methods, Eq. 9).

**a,b**, Green dots correspond to an example case of a region present at 6 copies, with 1 copy mutated, and purity 50%. Only 1 of 8 alleles at this locus carry the mutation (**a**), and 33 reads are required to detect the mutation with 80% power (**b**).

**c**, Theoretical power (as in **b**) to detect subclonal somatic point mutations present at multiplicity 1 in 0.2 of cancer cells.

**d,e,** Values were calculated for each tumor using the purity and ploidy estimates obtained from ABSOLUTE. Samples were considered powered if their detection power (as in **b**, **c**) exceeded 0.8.

**f**, Tumor purity and local absolute copy-number estimates were used to calculate the number of bases powered for the detection of subclonal mutations present at fraction 0.05 (*x*-axis, Online Methods, Eq. 9), and to compute the number of subclonal mutations at cancer-cell fraction <= 0.05 in each sample (*y*-axis). The dashed line,  $r^2$ , and *P*-value refer to a linear regression fit of the data points.



#### Supplementary figure 8 | Validation of subclonal mutations in ovarian cancer

a, b, Validation of somatic mutations in independent sequencing data. A total of 6050 of the 29,268 somatic mutations called in this study (a) were targeted for Illumina capture sequencing of genomic DNA from an independent whole genome amplification (WGA) reaction. Of these, 3901 were powered at ≥ 0.8 for detection in the validation sequencing (Eq. 9, Online Methods), of which 3713 (95%) were detected (b).
a, Number of clonal and subclonal mutations called above an allelic-fraction cutoff (note that all mutations were used in the analysis for Fig. 4).

**b**, Validation rate as a function of allelic fraction decile (x-axis; bin 1 = allelic fraction 0 to 0.1, etc).

c, d, Confirmation of sublclonal allelic fraction for validated mutations. A total of 316 putative subclonal mutations were detected in the validation data. Of these, 98 had power > 0.8 for discrimination of subclonal status

in the validation data (c), given the read depth (N), clonal allelic fraction ( $f_c$ ), and subclonal allelic fraction ( $H_a$ ,

as estimated in the original dataset). Of these 98, 83 were confirmed subclonal in the validation sequencing with strict Bonferroni corrected P-values < 0.05 / 316 (d).

c, Cumulative distribution of estimated discrimination power. Power was calculated by solving

$$\alpha = \int_0^y \operatorname{Beta}\left(f \mid f_c N + 1, (1 - f_c N) + 1\right) df \text{ for } y, \text{ and then}$$

 $\operatorname{Pow}(y, f_o, N, \alpha) = \int_0^{\infty} \operatorname{Beta}(f \mid f_o N + 1, (1 - f_o)N + 1) df$ , where  $\alpha$  was set to the significance level 0.05 / 316.

**d**, Cumulative distribution of  $-\log_{10} P$ -value for rejection of the clonal hypothesis. Dashed vertical line indicates the Bonferroni significance level 0.05 / 316.

b



### Supplementary figure 9 | Distribution of estimated ploidy in various tumor types, and correspondence with absolute homologous copy-numbers

**a,b**, Data are shown for each indicated sample-type, with samples sorted by cancer-genome ploidy (left). Absolute copy-numbers of genomic segments for each sample are partitioned into low (**a**) and high (**b**) homologous copy values. Genome doubling events are discernable as inflections in the ploidy distributions of each cancer (**a**; left), corresponding with an increase in genome-wide homologous copy-number (**a**,**b**).



## Supplementary figure 10 | Incidence of chromosome arm-level SCNAs in non genome-doubled (WGD=0) vs. genome doubled (WGD=1) samples from 10 cancer types.

As in **Fig. 6d**. LOH (loss of heterozygosity) was defined as 0 allelic copies. Amplification was defined as > 1 allelic copy for samples with 0 genome doublings, and as > 2 allelic copies for those with 1 genome doubling. Calls were made based on the modal allelic copy numbers of each chromosome arm. Dashed lines indicate y=x.



#### Supplementary figure 11 | Robustness of purity / ploidy estimation to experimental variability

**a**, Two estimates of cancer-genome ploidy:  $\hat{\tau}$  (eq. 1), and  $\hat{\tau}_g$  (eq. 6). The estimator  $\hat{\tau}$  is derived from the observed copy-ratio locations, whereas  $\hat{\tau}_g$  is derived from the expectation of discrete integer copy-states,

summed over the genome.

**b**, The discordance between the estimators in **a** (*y*-axis) is proportional to the mean of segmented copy-ratio data across the genome. In a perfectly calibrated experiment, this mean must equal 1, since a constant mass of DNA is input.

**c**, Correction of ploidy estimator  $\hat{\tau}$  by  $\hat{\tau} = (\hat{S} \cdot \hat{b} - 1)/(\hat{S} \cdot \hat{\delta}_{\tau})$  (x-axis), where  $\hat{\delta}_{\tau} = \hat{\alpha} / \hat{D}$ ,  $\hat{b} = 2(1 - \hat{\alpha}) / \hat{D}$ , with  $\hat{\alpha}$  and  $\hat{D}$  fit from the data (eqn. 1). This correction assumes that the data miscalibration is due to a *scale S*, which is estimated for each sample as 1 / the weighted mean of segmented copy-ratios (**b**; x-axis). Since this correction removes most of the discordance between  $\hat{\tau}$  and  $\hat{\tau}_g$ , we assume that the miscalibration in **b** is indeed primarily due to scaling variability. Noting that tumor purity  $\alpha = \delta_{\tau} / (\delta_{\tau} + b)$ , the implied correction for the purity estimate is:  $\hat{\alpha}' = \hat{S} \cdot \hat{\delta}_{\tau} / (\hat{S} \cdot \hat{\delta}_{\tau} + \hat{S} \cdot \hat{b}) = \hat{\alpha}$ . In other words, estimation of tumor purity is unaffected by moderate scaling variability in the data. This analysis also implies that centering of copy-ratio data, by rescaling (but not by translating) to mean = 1, is a valid preprocessing step (this was not performed). RMSE: root mean squared error



#### Supplementary figure 12 | Power for detection of subclonal somatic mutations in 8 HGS-OvCa samples.

Histograms of cellular multiplicity point-estimates are shown for individual tumor samples, in order of decreasing  $\delta_{\tau}$ , which determines the average detection power in each sample (**Supplementary Fig. 7e**). Individual point mutations are colored according to their estimated detection power, calculated using their observed allelic fraction, coverage, tumor purity, and local absolute copy-number (Online Methods, Eq. 9). Coverage values and absolute copy-numbers were used to calculate theoretical detection power at each base for subclonal fractions between 0.025 and 0.5 (Online Methods, Eq. 9). The number of bases with power >= 0.8 are shown for each fraction (blue curve, right axis).

### **Supplementary Note 1**

ASCAT was performed using version 2.1 downloaded

from http://heim.ifi.uio.no/bioinf/Projects/ASCAT/ on 01/29/2012. As is recommended for Affymetrix SNP data, the PennCNV-Affy protocol was used to convert probe-level data from CEL files to LRR and BAF values (http://www.ncbi.nlm.nih.gov/pmc/articles/PMC2045149/). A batch-specific canonical genotype cluster file was created via PennCNV-Affy steps 1.1 through 1.3 (http://www.openbioinformatics.org/penncnv/penncnv\_tutorial\_affy\_gw6.htm) using CEL files from all normal samples in that batch. Steps 1.1 and 1.2 were repeated on all CEL files, and step 1.4 was performed using the batch-specific canonical genotype cluster file to generate LRR and BAF values for the samples. Separate PennCNV-Affy protocols were used depending on the array type. The resulting LRR and BAF files were input into ASCAT as indicated on the software website.